



A LATENT VARIABLE REPRESENTATION OF COUNT DATA MODELS TO ACCOMMODATE SPATIAL AND TEMPORAL DEPENDENCE

**Application to Predicting Crash Frequency
at Intersections**

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Conteos

- Variable dependiente
 - Discreta
 - No negativa
 - Sin límite superior
- Ejemplos
 - Tasa de motorización
 - Viajes generados por hogar
 - Accidentes de tránsito



Modelos de Conteos

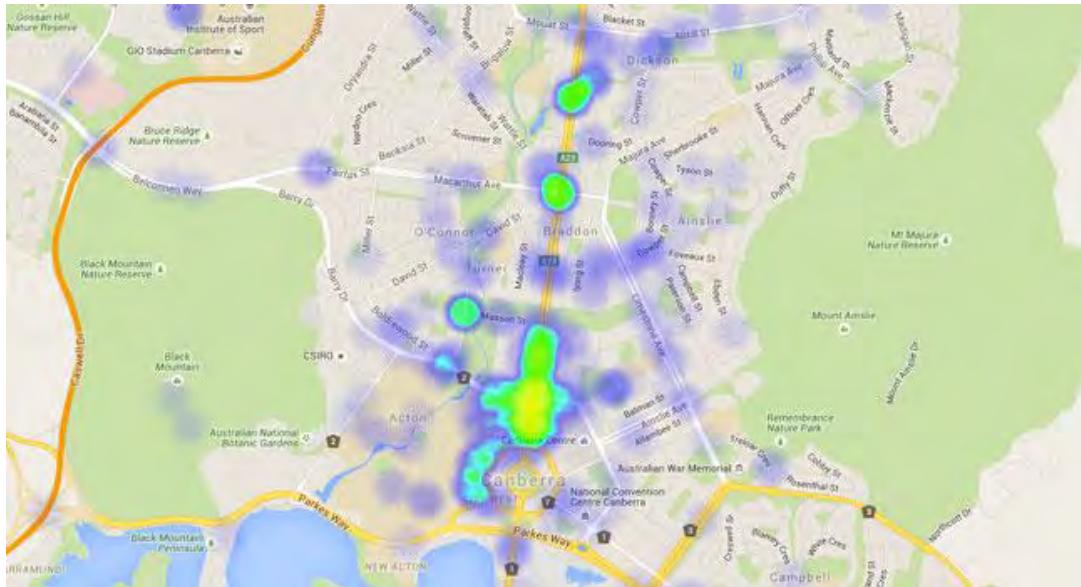
- Poisson
- Binomial negativa
- Binomial
- Logarítmica
- Zero-inflated, Hurdle

- Más usada para conteos
- Restricción varianza = media
- Permite sobredispersión
- **Modelos univariados:**
- Permite sub-dispersión
- Sin correlación temporal
- Problemas de estimación
- Sin correlación espacial
- Permite sobre y sub dispersión
- Problemas de estimación
- Controla el "exceso" de ceros

- **Modelos de conteos que sí incluyen correlaciones**
 - Poisson y binomial negativa multivariada
 - Mixture methods: términos aleatorios incluidos en la parametrización de la media
- Problemas con proceso de estimación: incluir correlación espacial y temporal es difícil/imposible

¿Correlación Temporal y Espacial?

- Correlación temporal
- Correlación espacial
 - Efectos no observados
 - Efectos de variables observadas (spillover effects)



Accidentes de ciclistas en Canberra

Alcance del Estudio

- Formular un nuevo modelo para conteos
 - Modelo ordinal
 - Incorpora estructuras flexibles de correlación espacial y temporal

MODELO

Modelo Ordinal

$$y_q^* = \boldsymbol{\beta}' \mathbf{x}_q + \varepsilon_q$$

$$y_q = k \quad \text{if} \quad \psi_{k-1} < y_q^* < \psi_k$$

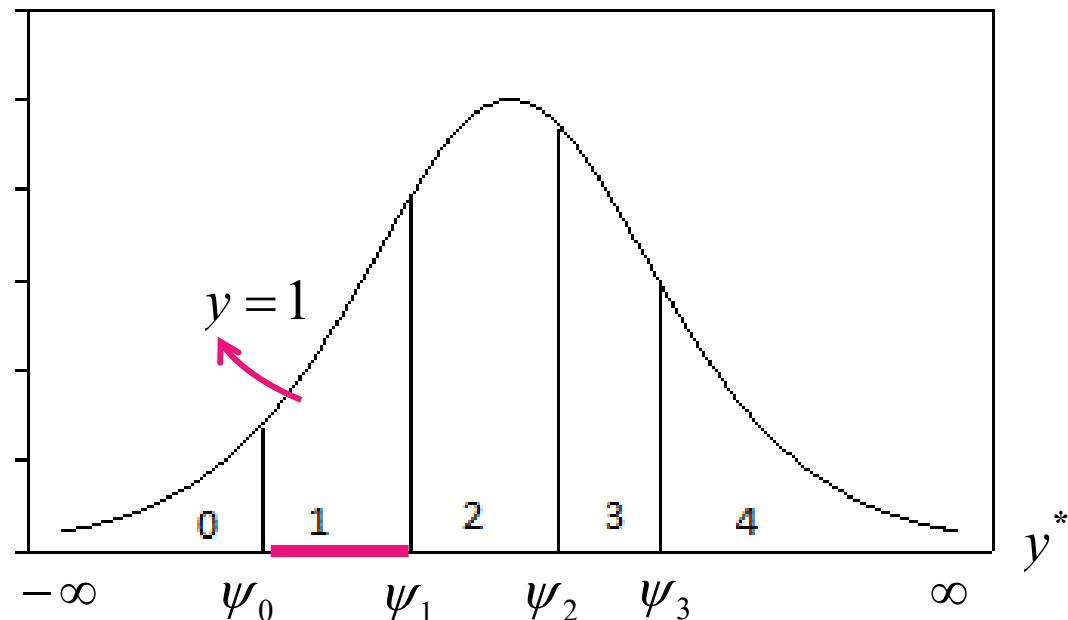
Variable latente

$q = 1, 2, \dots, N$
unidad de análisis

Variable dependiente

$k = 0, 1, \dots, K$
Nivel (ordinal)

Umbrales $-\infty < \psi_1 < \psi_2 < \dots < \psi_{k-1} < \infty$



Modelo Ordinal

$$y_q^* = \beta' x_q + \varepsilon_q$$

$$y_q = k \quad \text{if} \quad \psi_{k-1} < y_q^* < \psi_k$$

Variable latente

$q = 1, 2, \dots, N$
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Nivel (ordinal)

Umbrales $-\infty < \psi_1 < \psi_2 < \dots < \psi_{k-1} < \infty$

- Parametrización de los umbrales $\psi_{qk} = f_k(z_q)$
 - Umbrales (mapping) puede variar entre observaciones

Φ^{-1} función inversa de la distribución normal univariada

α_k constante asociada al umbral k

γ parámetro a estimar

z_q variables asociadas al umbral

$$\psi_{qk} = \Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) + \alpha_k$$

donde $\lambda_q = e^{\gamma' z_q}$

Incluyendo correlaciones...

- Correlación es incorporada a través de la variable latente y^*
- Modelo

$$y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta_q' x_{qt} + \varepsilon_{qt}$$

$q = 1, 2, \dots, N$ unidad de análisis
 $t = 1, 2, \dots, T$ periodo de tiempo

- $w_{qq'}$ matriz espacial (weight matrix) $\rightarrow W (N \times N)$
- δ parámetro espacial (escalar)
- $\beta_q \sim N(\mathbf{b}, \Omega)$ parámetro aleatorio, $(L \times 1) \rightarrow \beta_q = \mathbf{b} + \tilde{\beta}_q$, $\tilde{\beta}_q \sim MVN_L(0, \Omega)$
- $\varepsilon_{qt} \sim N(0, \Lambda)$ término de error

$$\Lambda = \begin{bmatrix} 1 & \rho & \rho^2 & \rho^3 & \dots & \rho^{T-1} \\ \rho & 1 & \rho & \rho^2 & \dots & \rho^{T-2} \\ \rho^2 & \rho & 1 & \rho & \dots & \rho^{T-3} \\ \vdots & \vdots & \vdots & 1 & \dots & \vdots \\ \rho^{T-1} & \rho^{T-2} & \rho^{T-3} & \rho^{T-4} & \dots & 1 \end{bmatrix}$$

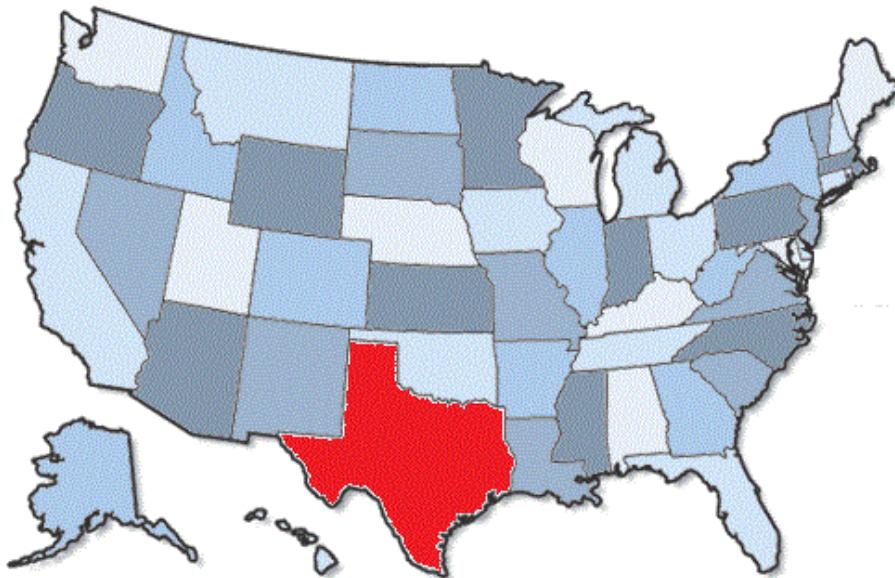
Estructura temporal autoregresiva

$\Lambda: (T \times T)$

APLICACIÓN DEL MODELO

Análisis de accidentes en intersecciones

Base de Datos

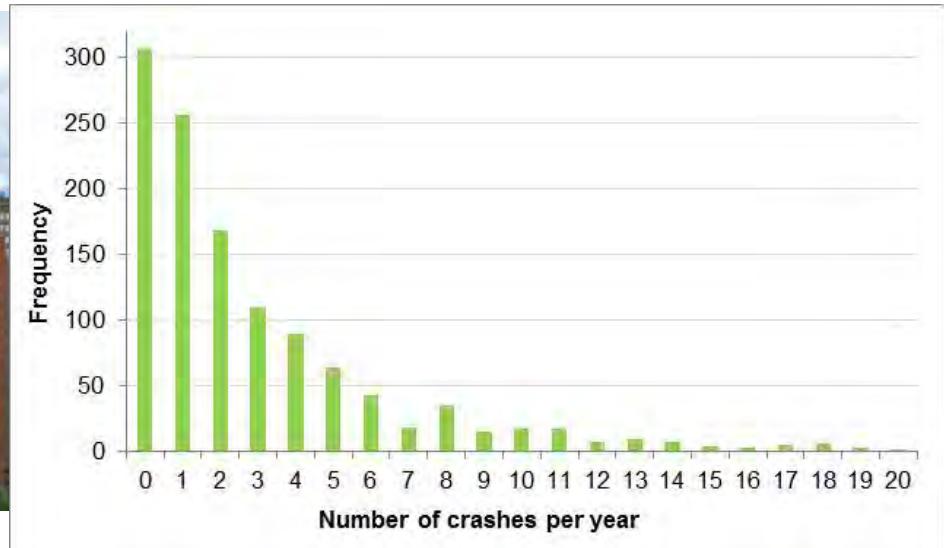


Estado of Texas



Arlington

- 170 intersecciones
- 7 años
- 1190 observaciones
- 3,503 accidentes
 - Promedio 2,94 por año-intersección



Variables	No Correlation		Spatial and Temporal Effects	
	Estimate	t-stat	Estimate	t-stat
Threshold Variables				
Threshold Specific Constants				
α_1	-0.026	-1.06	0.492	10.30
α_2	-0.166	-5.02	0.708	10.16
α_3	-0.377	-8.81	0.769	8.77
α_4	-0.563	-11.25	0.842	7.93
α_5	-0.745	-13.02	0.901	7.23
α_6	-0.968	-15.18	0.882	6.20
α_7	-1.264	-17.25	0.722	4.65
α_8	-1.430	-18.88	0.809	4.65
α_9	-1.698	-20.81	0.689	3.70
γ - Vector				
Constant	1.223	40.40	2.309	16.09
<i>Approach Roadway Type</i>				
At least one approach roadway is non-city street	0.449	11.39	0.313	9.00
<i>Type of Traffic Control</i>				
Regular signal light	0.837	7.49	0.888	10.82

$$\psi_{q,m_{qt},t} = \Phi^{-1} \left(e^{-\lambda_{qt}} \sum_{l=0}^{m_{qt}} \frac{\lambda_{qt}^l}{l!} \right) + \alpha_{k_{qt}}$$

$$\lambda_{qt} = e^{\gamma' z_{qt}}$$

$$\alpha_0 = 0, \quad \alpha_{k_{qt}} = \alpha_K \text{ if } k_{qt} > K,$$

$$y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta_q' x_{qt} + \varepsilon_{qt},$$

$\delta = 0.422$ (9.25)
(parámetro de correlación espacial)

Variables	No Correlation		Spatial and Temporal Effects	
	Estimate	t-stat	Estimate	t-stat
Latent Variables				
Constant	0.000	-	0.000	-
Standard Deviation	-	-	0.640	7.02
Number of Entering Roads				
Three	-0.660	-15.07	-0.950	-13.20
More than four	-0.921	-13.64	-1.276	-13.20
Type of Traffic Control				
Regular signal light	-1.906	-7.35	-3.003	-7.53
Yield sign	-0.912	-10.40	-1.352	-8.55
Standard Deviation	-	-	1.168	6.32
Stop sign	-0.454	-7.80	-0.540	-7.01
Flashing light	0.696	8.69	0.948	9.08
Center stripe / divider	-0.659	-8.83	-0.786	-6.96
Logarithm of daily entering volume (veh/day/10.000)	0.210	6.76	0.374	7.98
Standard Deviation	-	-	0.831	15.60
Flow split imbalance (FSIMB) factor	-0.817	-9.86	-1.042	-9.49
Year-specific Dummy Variables				
Year 2004	-0.210	-2.23	-0.103	-1.09
Year 2005	-0.440	-4.67	-0.291	-2.90
Year 2006	-0.438	-4.69	-0.218	-1.94
Year 2007	-0.441	-4.67	-0.182	-1.52
Year 2008	-0.470	-5.03	-0.250	-2.10
Year 2009	-0.539	-5.85	-0.378	-3.30

Comentarios finales

- Modelo aplicable a distintos casos de estudio
- (Potencial) importancia de incluir estructuras de correlación
- En el paper hay detalles de:
 - Interpretación del modelo y sus parámetros
 - Método de estimación
 - Elección de la matriz espacial
 - Resultados de la aplicación

GRACIAS!!

Castro, M., Paleti, R., & Bhat, C. R. (2012).

A latent variable representation of count data models to accommodate spatial and temporal dependence: Application to predicting crash frequency at intersections.

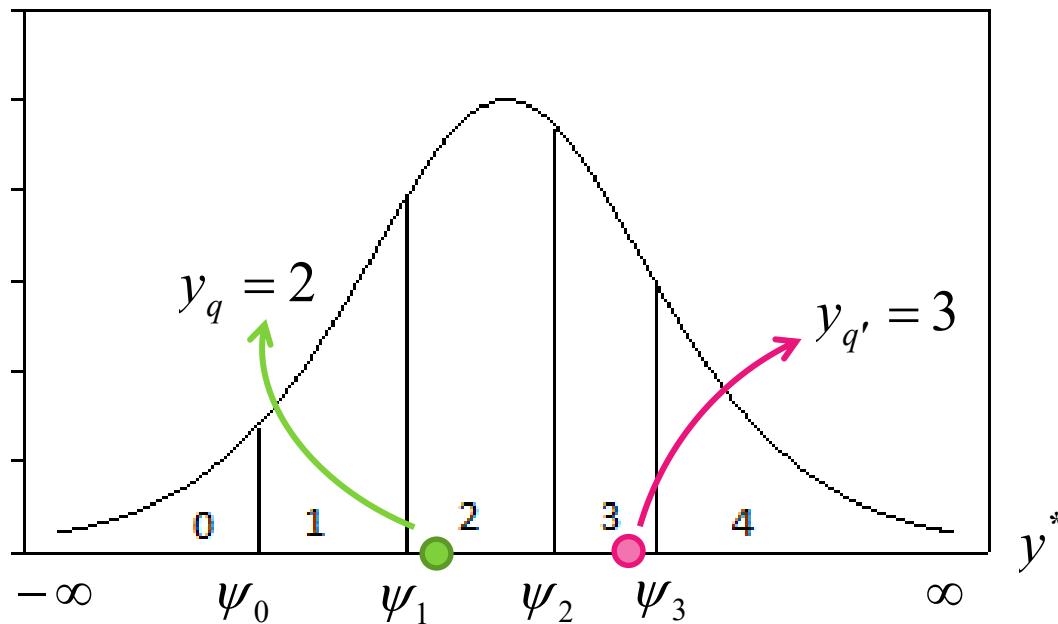
Transportation research part B: methodological, 46(1), 253-272.

How do we interpret this representation?

$$\begin{aligned} y_q^* &= \boldsymbol{\beta}' \mathbf{x}_q + \varepsilon_q \\ y_q = k &\quad \text{if } \psi_{q,k-1} < y_q^* < \psi_{q,k} \end{aligned}$$

$$\psi_{qk} = \Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) + \alpha_k \quad \text{where} \quad \lambda_q = e^{\gamma' z_q}$$

- Two sources of variability
 - Latent variable
 - Threshold

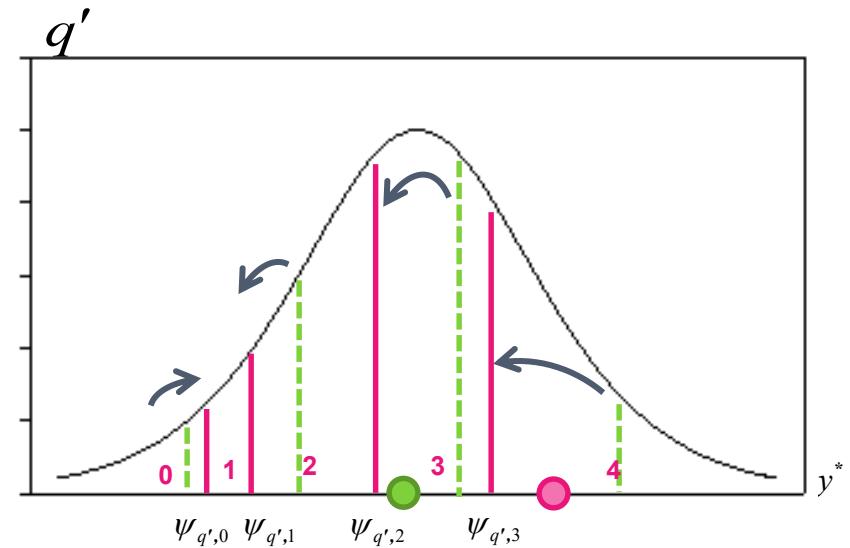
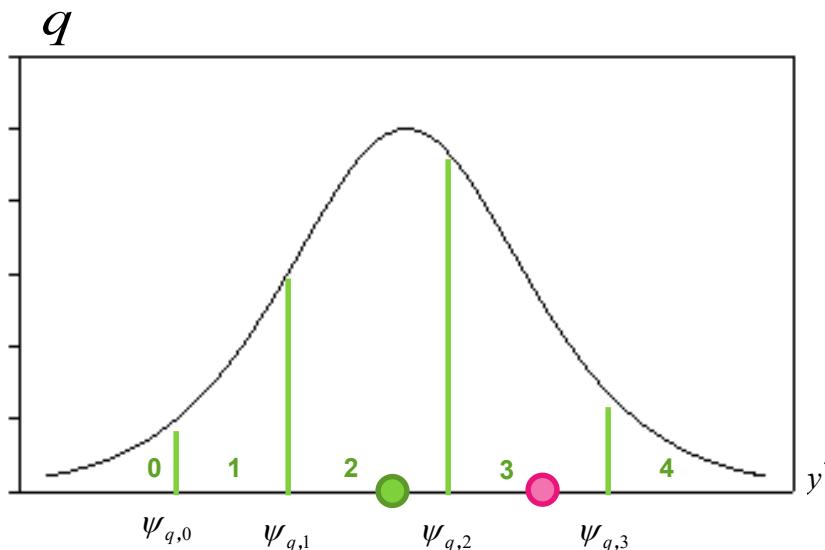


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- Two sources of variability
 - Latent variable \rightarrow latent long-term propensity
 - Threshold \rightarrow instantaneous mapping



Why choose this parameterization?

$$y_q^* = \boldsymbol{\beta}' \mathbf{x}_q + \varepsilon_q$$

$$y_q = k \quad \text{if} \quad \psi_{q,k-1} < y_q^* < \psi_{q,k}$$

$$\psi_{qk} = \Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) + \alpha_k \quad \text{where} \quad \lambda_q = e^{\gamma' z_q}$$

1. Satisfy threshold ordering condition

2. Allows identification for variables common in both x and z

3. Is a generalization of the Poisson model

- $\alpha_k = 0$
- $\beta = 0$
- $\varepsilon \sim N(0,1)$

$$P[y_q = k] = P \left[\Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^{k-1} \frac{\lambda_q^l}{l!} \right) < y_q^* < \Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) \right]$$

$$= \Phi \left(\Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) \right) - \Phi \left(\Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^{k-1} \frac{\lambda_q^l}{l!} \right) \right)$$

$$= \frac{e^{-\lambda_q} \lambda_q^k}{k!}$$

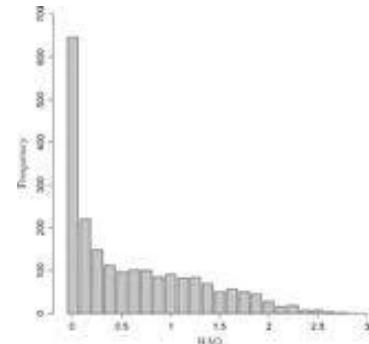
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$$y_q = k \quad \text{if} \quad \psi_{q,k-1} < y_q^* < \psi_{q,k}$$

$$\psi_{qk} = \Phi^{-1} \left(e^{-\lambda_q} \sum_{l=0}^k \frac{\lambda_q^l}{l!} \right) + \alpha_k \quad \text{where} \quad \lambda_q = e^{\gamma' z_q}$$

1. Satisfy threshold ordering condition
2. Allows identification for variables common in both x and z
3. Is a generalization of the Poisson model
4. Data does not require an upper bound K
5. The threshold-specific constants α can assign different probability mass to the counts
6. Can handle flexible correlation structures



Special Cases

Latent variable representation

$$y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta_q' x_{qt} + \varepsilon_{qt},$$

$$\beta \sim N(0, \Omega)$$

$$y_{qt} = k_{qt} \quad \text{if} \quad \psi_{q,k_{qt}-1,t} < y_{qt}^* < \psi_{q,k_{qt},t}$$

Threshold parameterization

$$\psi_{q,m_{qt},t} = \Phi^{-1} \left(e^{-\lambda_{qt}} \sum_{l=0}^{m_{qt}} \frac{\lambda_{qt}^l}{l!} \right) + \alpha_{k_{qt}}$$

$$\lambda_{qt} = e^{\gamma' z_{qt}}$$

$$\alpha_0 = 0, \quad \alpha_{k_{qt}} = \alpha_K \text{ if } k_{qt} > K,$$

- $\rho = 0$ Lack of time-varying temporal correlation
- $\delta = 0$ No spatial correlation
- non-diagonal elements of $\Omega = 0$ (diagonal elements non-zero)
Presence of time-invariant and unobserved heterogeneity effects, but without correlation between these effects
- $\Omega = 0$ Lack of time-invariance and unobserved heterogeneity
- $\rho = \delta = \Omega = 0$ Ordered response model without correlation
- $\rho = \delta = \Omega = \alpha = 0$ Poisson model (without correlation)

Estimation

$$L(\boldsymbol{\theta}) = P(\mathbf{y} = \mathbf{m}) = \int_{D_{\mathbf{y}^*}} \phi_{QT}(\mathbf{y}^* | \mathbf{b}, \Sigma) d\mathbf{y}^*,$$

$D_{\mathbf{y}^*}$: multivariate region of the elements of \mathbf{y}^*
 $\boldsymbol{\theta} = (\mathbf{b}', \boldsymbol{\Omega}', \rho, \delta, \gamma', \alpha')'$
 $(QT \times QT)$

- Composite marginal likelihood (CML)

- Used when the likelihood function is difficult/impossible to evaluate
- Estimator:
 - is consistent and asymptotically normal distributed
 - loses some asymptotic efficiency relative to a likelihood estimator
- Maximizes a surrogate likelihood function that compounds much easier-to-compute, lower-dimensional, marginal likelihoods
 - Pairwise CML: bivariate cumulative normal distribution

$$L_{CML}(\boldsymbol{\theta}) = \left(\prod_{g=1}^{QT-1} \prod_{g'=g+1}^{QT} P([\mathbf{y}]_g = [\mathbf{k}]_g, [\mathbf{y}]_{g'} = [\mathbf{k}]_{g'}) \right)^{(2 \times 2)} \quad \frac{QT(QT-1)}{2} \text{ times}$$

- Standard Errors $V_{CML}(\hat{\boldsymbol{\theta}}) = [\mathbf{G}(\hat{\boldsymbol{\theta}})]^{-1} = [\mathbf{H}(\hat{\boldsymbol{\theta}})]^{-1} \mathbf{J}(\hat{\boldsymbol{\theta}}) [\mathbf{H}(\hat{\boldsymbol{\theta}})]^{-1}$

$$\hat{\mathbf{H}}(\hat{\boldsymbol{\theta}}) = - \left[\sum_{g=1}^{QT-1} \sum_{g'=g+1}^{QT} \frac{\partial^2 \log L_{CML,gg'}(\boldsymbol{\theta})}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \right]_{\hat{\boldsymbol{\theta}}} \quad \mathbf{J}(\hat{\boldsymbol{\theta}}) = \frac{\tilde{W}}{D} \left[\sum_{d=1}^D \left[\frac{1}{N_d} ([\mathbf{s}_{CML,d}(\boldsymbol{\theta})][\mathbf{s}_{CML,d}(\boldsymbol{\theta})']_{\hat{\boldsymbol{\theta}}} \right] \right]$$

Estimation

- Spatial correlation $y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta_q' x_{qt} + \varepsilon_{qt}$

- Spatial weight matrix

- $1/d_{ij}^p$ $p = 1, 2, 3, \dots$
- $1/\exp(d_{ij})$
- 1 if $d_{ij} < U$

- Distance bands

- Explore different bands
- Minimize the variance $\text{tr}[V_{CML}(\hat{\theta})]$
- 2, 3, 4, 5, 6, 7, and 11.81 miles

- Models

- No correlation, $\rho = \delta = \Omega = 0$
- Temporal correlation, $\rho, \Omega \neq 0, \delta = 0$
- Temporal and spatial correlation, $\rho, \Omega, \delta \neq 0$

- Best model ?
- $\text{ADCLR} \sim \chi^2$

ADCLRT:
1646

→ 27 parameters
→ 30 parameters
→ 31 parameters



ADCLRT:
356

Estimation

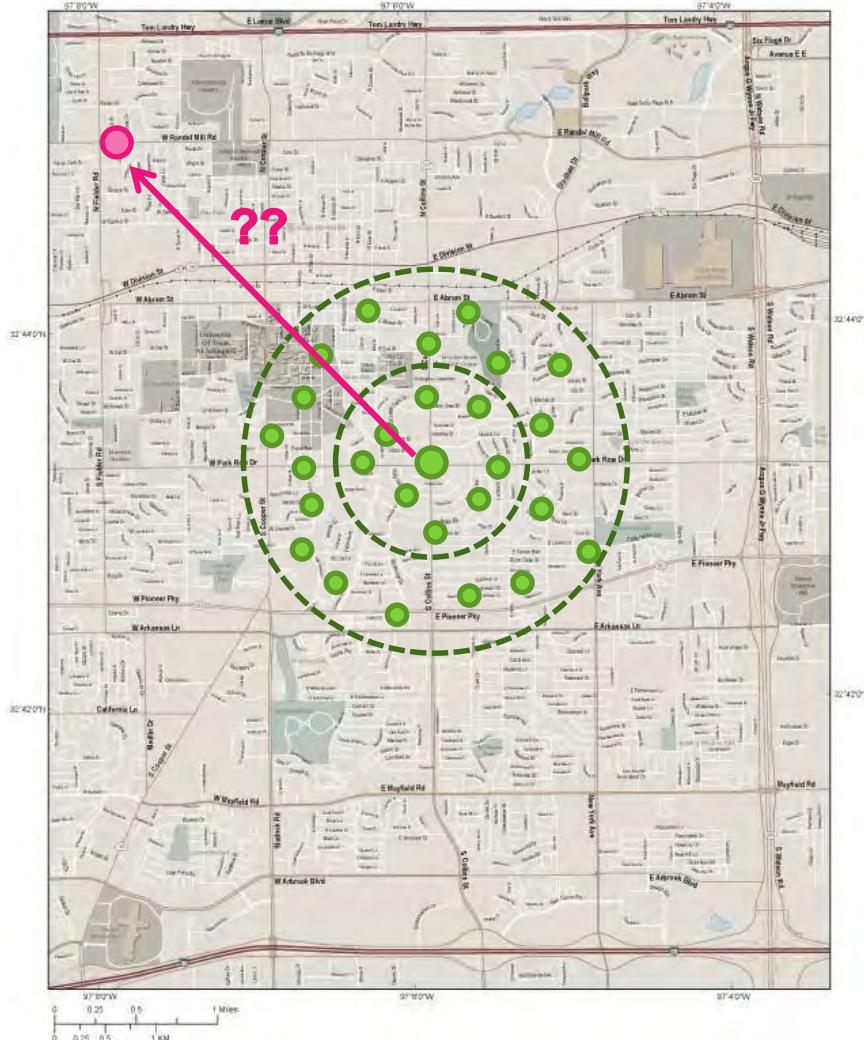
- Spatial correlation $y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta'_q x_{qt} + \varepsilon_{qt}$

- Spatial weight matrix

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- Spatial correlation $y_{qt}^* = \delta \sum_{q'=1}^Q w_{qq'} y_{q't}^* + \beta_q' x_{qt} + \varepsilon_{qt}$

- Spatial weight matrix

- $1/d_{ij}^p$ $p = 1, 2, 3, \dots$
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- Distance bands

- Explore different bands
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- No correlation, $\rho = \delta = \Omega = 0$
- Temporal correlation, $\rho, \Omega \neq 0, \delta = 0$
- Temporal and spatial correlation, $\rho, \Omega, \delta \neq 0$

- Best model ?

- ADCLR $\sim \chi^2$

ρ not significant

Variable	No correlation		Spatial and Temporal Effects	
	Estimate	Standard error	Estimate	Standard error
<i>Number of Entering Roads</i>				
Three	-51.07	4.31	-54.44	5.63
More than four	-60.95	4.07	-75.08	4.39
<i>Approach Roadway Type Combination</i>				
At least one approach roadway is a non-city street	133.44	20.20	85.22	15.63
<i>Type of Traffic Control</i>				
Regular signal light	-2.85	33.24	-7.87	36.31
Yield sign	-59.45	5.04	-69.50	6.78
Stop sign	-34.79	4.78	-37.50	5.83
Flashing light	69.45	11.94	99.40	21.43
Center stripe / divider	-47.36	5.53	-53.71	7.42
Logarithm of daily entering volume (veh/day/10,000) – 10% increase	1.75	0.31	5.06	1.09
Flow split imbalance (FSIMB) factor – 0.1 increase	-6.88	0.98	-9.74	1.65

Variables	Sample share			
<i>Number of Entering Roads</i>				
Three	24.6			
Four	71.6			
More than four	3.8			
<i>Roadway Alignment</i>				
All approaches are straight with no vertical grade	95.7			
At least one approach has horizontal curvature or vertical grade	4.3			
<i>Approach Roadway Type Combination</i>				
All approach roadways are city streets	94.4			
At least one approach roadway is a non-city street	5.6			
<i>Type of Traffic Control</i>				
Regular signal light	52.8			
Yield sign	15.3			
Stop sign	12.4			
Flashing light	7.3			
Center stripe / divider	4.6			
No traffic control or minimal traffic control	7.6			
Descriptive Statistics				
	Minimum	Maximum	Mean	Std. Dev.
Total daily entering volume (vehicles/day)	2,866	193,178	35,222	33,784
Flow split imbalance (FSIMB) factor	0.00	0.97	0.43	0.25
Distance between intersections (miles)	0.05	11.81	4.42	2.29